

Sparse Additive Interaction Learning

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<https://sahirbhatnagar.com/sail>
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Outline

Betting on Sparsity

A Thought Experiment

Motivating Example: The Nurse Family Partnership

sail: Strong Additive Interaction Learning

Algorithm

Theory

Simulations

sail R package

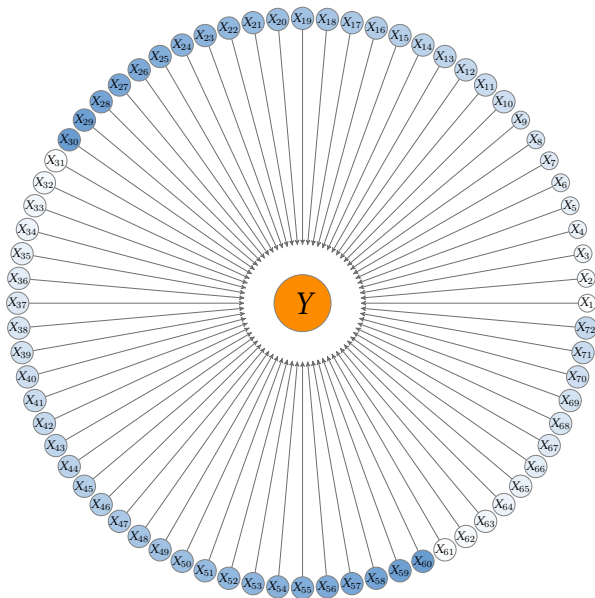
Real Data Application

Discussion

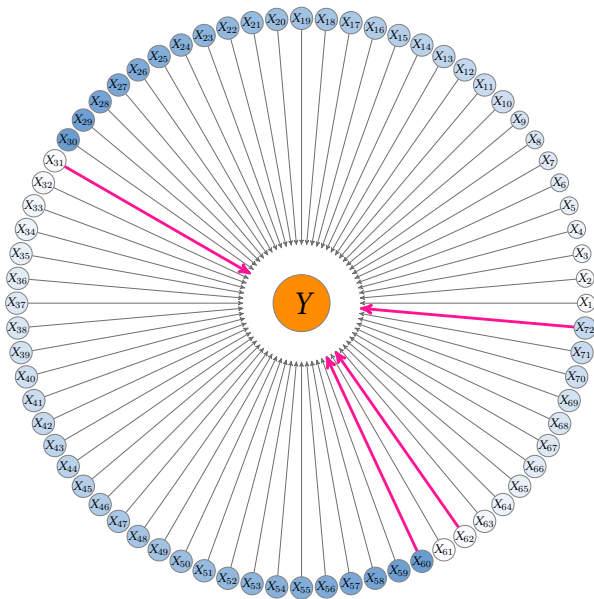
Current and Future Work

Acknowledgements

Bet on Sparsity Principle



Bet on Sparsity Principle



Bet on Sparsity Principle

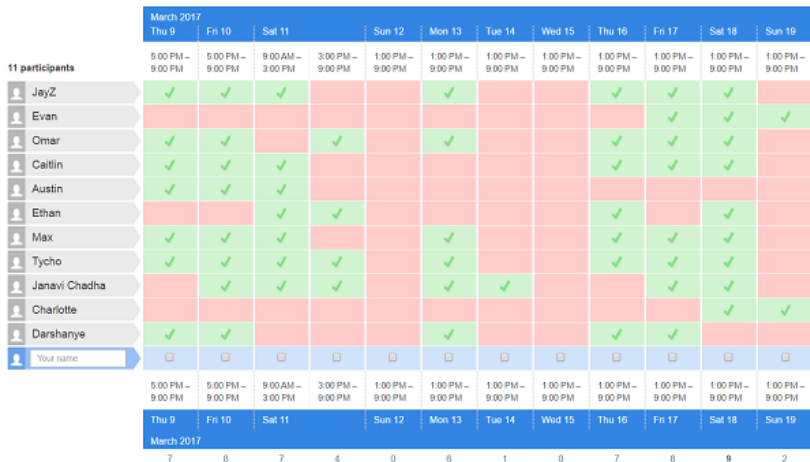
Use a procedure that does well in sparse problems, since no procedure does well in dense problems.¹

- We often don't have enough data to estimate so many parameters
- Even when we do, we might want to identify a **relatively small number of predictors** ($k < N$) that play an important role
- Faster computation, easier to understand, and stable predictions on new datasets.

¹The elements of statistical learning. Springer series in statistics, 2001.

How would you schedule a meeting of 20 people?

How would you schedule a meeting of 20 people?



Doctors Bet on Sparsity Also

Doctors Bet on Sparsity Also

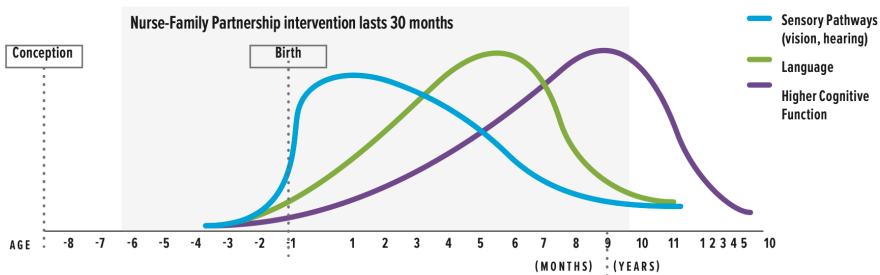




Nurse-Family Partnership is an evidence-based, community health program with over 40 years of evidence showing significant improvements in the health and lives of first-time moms and their children living in poverty.

Human Brain Development

Synapse formation dependent on early experiences



Source: Nelson, C.A., In *Neurons to Neighborhoods* (2000).



ALL CAUSE MORTALITY OVER 20 YEAR FOLLOW-UP

3x

Mothers who did not receive nurse home visits were nearly **3 times more likely to die** from all causes of death than nurse visited mothers (3.7% versus 1.3%)¹

8x

Mothers that did not receive nurse home visits were **8 times more likely to die** from external causes – including unintentional injuries, suicide, drug overdose and homicide – than nurse visited mothers (1.7% versus 0.2%)¹



PREVENTABLE CHILD MORTALITY OVER 20 YEAR FOLLOW-UP

- Among Nurse-Family Partnership participants, there were **lower rates of preventable child mortality** from birth until age 20.¹
- 1.6% of the children not receiving nurse home visits died from preventable causes – including sudden infant death syndrome, unintentional injuries and homicide – while none of the nurse visited children died from these causes.¹

Additional Maternal and Child Health Outcomes

Maternal Health Outcomes

35% fewer cases of pregnancy-induced hypertension⁵

18% fewer preterm births⁶

79% reduction in preterm delivery among women who smoke cigarettes⁷

31% reduction in very closely spaced (<6 months) subsequent pregnancies⁸

Child Health Outcomes

48% reduction in child abuse and neglect⁹

39% fewer health care encounters for injuries or ingestions in the first 2 years of life among children born to mothers with low psychological resources¹⁰

67% less behavioral and intellectual problems in children at age 6¹¹

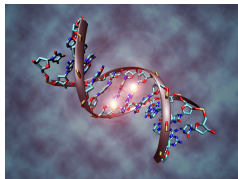
56% fewer emergency room visits for accidents and poisonings through age 2¹²

Interactions between Intervention and Genetics

Stanford-Binet Fifth Edition (SB5) classification^[4]

| IQ Range ("deviation IQ") | IQ Classification |
|---------------------------|--------------------------------|
| 145–160 | Very gifted or highly advanced |
| 130–144 | Gifted or very advanced |
| 120–129 | Superior |
| 110–119 | High average |
| 90–109 | Average |
| 80–89 | Low average |
| 70–79 | Borderline impaired or delayed |
| 55–69 | Mildly impaired or delayed |
| 40–54 | Moderately impaired or delayed |

~



×



Phenotype
IQ Score

Large Data
Genetic Markers

Environment
NFP Intervention

$$Y = \sum_{j=1}^p X_j \beta_j + \sum_{j=1}^p X_j X_E \tau_j + \varepsilon$$

$$\begin{aligned}
Y &= \sum_{j=1}^p X_j \beta_j & + & \sum_{j=1}^p X_j X_E \tau_j & + & \varepsilon \\
= & \begin{array}{c} \text{[Blue } n \times p \text{ matrix } \mathbf{X}] \text{ [Yellow } p \times 1 \text{ vector } \boldsymbol{\beta}] \\ n \times p \qquad p \times 1 \end{array} & + & \begin{array}{c} \text{[Red } n \times 1 \text{ vector } X_E] \circ \text{ [Blue } n \times p \text{ matrix } \mathbf{X}] \text{ [Green } p \times 1 \text{ vector } \boldsymbol{\tau}] \\ n \times 1 \qquad n \times p \qquad p \times 1 \end{array} & + & \begin{array}{c} \text{[Grey } n \times 1 \text{ vector } \boldsymbol{\varepsilon}] \\ n \times 1 \end{array}
\end{aligned}$$

$$\begin{aligned}
 Y &= \sum_{j=1}^p X_j \beta_j & + & \sum_{j=1}^p X_j X_E \tau_j & + & \varepsilon \\
 &= & & & & \\
 & \begin{array}{c} \text{Main effects} \end{array} & + & \begin{array}{c} \text{Interaction effects} \end{array} & + & \begin{array}{c} \text{Error} \end{array}
 \end{aligned}$$

The diagram illustrates the matrix representation of a linear model with main and interaction effects. The response variable Y is an $n \times 1$ vector. The design matrix is composed of three parts:

- Main effects:** A blue square matrix \mathbf{X} of size $n \times p$ multiplied by a yellow vector $\boldsymbol{\beta}$ of size $p \times 1$.
- Interaction effects:** A red vector X_E of size $n \times 1$ (representing the interaction term) multiplied by a blue square matrix \mathbf{X} of size $n \times p$ (representing the main effects), resulting in a green vector $\boldsymbol{\tau}$ of size $p \times 1$.
- Error:** A gray vector $\boldsymbol{\varepsilon}$ of size $n \times 1$.

The dimensions of the matrices and vectors are indicated below them:

- \mathbf{X} (Main effects): $n \times p$
- $\boldsymbol{\beta}$ (Main effects): $p \times 1$
- X_E (Interaction effects): $n \times 1$
- \mathbf{X} (Interaction effects): $n \times p$
- $\boldsymbol{\tau}$ (Interaction effects): $p \times 1$
- $\boldsymbol{\varepsilon}$ (Error): $n \times 1$

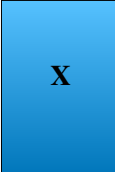
$$\begin{aligned}
 Y &= \sum_{j=1}^p X_j \beta_j + \sum_{j=1}^p X_j X_E \tau_j + \varepsilon \\
 &= \begin{array}{c} \text{Main effects} \end{array} + \begin{array}{c} \text{Interaction effects} \end{array} + \begin{array}{c} \text{Error} \end{array}
 \end{aligned}$$

Diagram illustrating the components of a linear model:


- Main effects:** A blue square matrix \mathbf{X} of size $n \times p$ multiplied by a yellow column vector $\boldsymbol{\beta}$ of size $p \times 1$.
- Interaction effects:** A red column vector X_E of size $n \times 1$ multiplied by a blue square matrix \mathbf{X} of size $n \times p$, resulting in a green column vector $\boldsymbol{\tau}$ of size $p \times 1$.
- Error:** A gray column vector $\boldsymbol{\varepsilon}$ of size $n \times 1$.

Let $Z_{jE} = X_E X_j$


$$\begin{aligned}
 Y &= \sum_{j=1}^p X_j \beta_j + \sum_{j=1}^p Z_{jE} \tau_j + \epsilon \\
 &= \begin{array}{|c|c|} \hline \mathbf{X} & \mathbf{Z} \\ \hline \end{array} \begin{array}{c} \beta \\ \tau \end{array} + \epsilon
 \end{aligned}$$




$n \times p$




$n \times p$



$n \times 1$

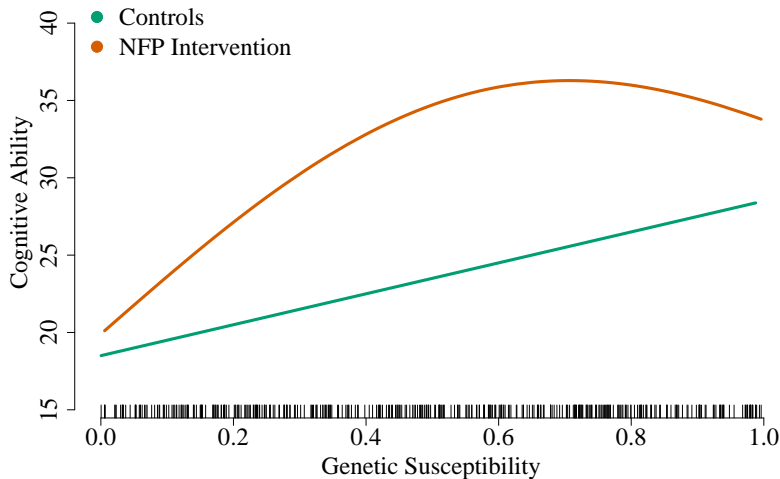


$2p \times 1$



$n \times 1$

Motivation 1: Non-linear Interactions



Motivation 2: Heredity Property

$$Y = \beta_0 \cdot \mathbf{1} + \underbrace{\sum_{j=1}^p \beta_j X_j + \beta_E X_E}_{\text{main effects}} + \underbrace{\sum_{j=1}^p \tau_j X_E X_j}_{\text{interactions}} + \varepsilon$$

¹Chipman. Canadian Journal of Statistics (1996)

²McCullagh and Nelder. Generalized Linear Models (1983)

³Cox. International Statistical Review (1984)

Motivation 2: Heredity Property

$$Y = \beta_0 \cdot \mathbf{1} + \underbrace{\sum_{j=1}^p \beta_j X_j + \beta_E X_E}_{\text{main effects}} + \underbrace{\sum_{j=1}^p \tau_j X_E X_j}_{\text{interactions}} + \varepsilon$$

Strong Heredity¹

$$\hat{\tau}_j \neq 0 \quad \Rightarrow \quad \hat{\beta}_j \neq 0 \quad \text{and} \quad \hat{\beta}_E \neq 0$$

- Heredity property is desired for the purposes of **interpretability**²
- Large main effects are more likely to lead to appreciable interactions³

¹Chipman. Canadian Journal of Statistics (1996)

²McCullagh and Nelder. Generalized Linear Models (1983)

³Cox. International Statistical Review (1984)

Lasso interaction model

- $Y \rightarrow$ response
- $X_E \rightarrow$ environment
- $X_j \rightarrow$ predictors, $j = 1, \dots, p$

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \beta_j X_j + \beta_E X_E + \sum_{j=1}^p \tau_j X_E X_j + \varepsilon$$

$$\underset{\Theta := (\beta_0, \boldsymbol{\beta}, \boldsymbol{\tau})}{\operatorname{argmin}} \quad \mathcal{L}(\Theta) + \lambda(\|\boldsymbol{\beta}\|_1 + \|\boldsymbol{\tau}\|_1)$$

Strong Heredity Interactions: Current State of the Art

| Type | Model | Software |
|------------|--|--------------------------------------|
| Linear | CAP (Zhao et al. 2009, Ann. Stat) | X |
| | SHIM (Choi et al. 2009, JASA) | X |
| | hiernet (Bien et al. 2013, Ann. Stat) | hierNet(x, y) |
| | GRESH (She and Jiang 2014, JASA) | X |
| | FAMILY (Haris et al. 2014, JCGS) | FAMILY(x, z, y) |
| | glinternet (Lim and Hastie 2015, JCGS) | glinternet(x, y) |
| | RAMP (Hao et al. 2016, JASA) | RAMP(x, y) |
| | LassoBacktracking (Shah 2018, JMLR) | LassoBT(x, y) |
| Non-linear | VANISH (Radchenko and James 2010, JASA) | X |
| | sail (Bhatnagar et al. 2020+, in revision CSDA) | sail(x, e, y, basis) |

Our Extension to Nonlinear Effects

Consider the basis expansion

$$f_j(X_j) = \sum_{\ell=1}^{m_j} \psi_{j\ell}(X_j) \beta_{j\ell}$$

$$f(X_1) = \underbrace{\begin{bmatrix} \psi_{11}(X_{11}) & \psi_{12}(X_{12}) & \cdots & \psi_{11}(X_{15}) \\ \vdots & \vdots & \cdots & \vdots \\ \vdots & \vdots & \cdots & \vdots \\ \psi_{11}(X_{i1}) & \psi_{12}(X_{i2}) & \cdots & \psi_{11}(X_{i5}) \\ \vdots & \vdots & \cdots & \vdots \\ \vdots & \vdots & \cdots & \vdots \\ \vdots & \vdots & \cdots & \vdots \\ \psi_{11}(X_{N1}) & \psi_{12}(X_{N2}) & \cdots & \psi_{11}(X_{N5}) \end{bmatrix}}_{\Psi_1} \quad N \times 5 \quad \times \quad \underbrace{\begin{bmatrix} \beta_{11} \\ \beta_{12} \\ \beta_{13} \\ \beta_{14} \\ \beta_{15} \end{bmatrix}}_{\theta_1} \quad 5 \times 1$$

sail: Additive Interactions

- $\boldsymbol{\theta}_j = (\beta_{j1}, \dots, \beta_{jm_j}) \in \mathbb{R}^{m_j}$
- $\boldsymbol{\tau}_j = (\tau_{j1}, \dots, \tau_{jm_j}) \in \mathbb{R}^{m_j}$
- $\boldsymbol{\Psi}_j \rightarrow n \times m_j$ matrix of evaluations of the $\psi_{j\ell}$
- In our implementation, we use cubic bsplines with 5 degrees of freedom

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \boldsymbol{\Psi}_j \boldsymbol{\theta}_j + \beta_E X_E + \sum_{j=1}^p (X_E \circ \boldsymbol{\Psi}_j) \boldsymbol{\tau}_j + \varepsilon$$

sail: Strong Heredity

Reparametrization¹

$$\boldsymbol{\tau}_j = \gamma_j \beta_E \boldsymbol{\theta}_j$$

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \boldsymbol{\Psi}_j \boldsymbol{\theta}_j + \beta_E X_E + \sum_{j=1}^p \gamma_j \beta_E (X_E \circ \boldsymbol{\Psi}_j) \boldsymbol{\theta}_j + \varepsilon$$

Objective Function

$$\underset{\boldsymbol{\Theta} := (\beta_E, \boldsymbol{\theta}, \boldsymbol{\gamma})}{\operatorname{argmin}} \quad \mathcal{L}(\boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\boldsymbol{\theta}_j\|_2 \right) + \lambda \alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

¹Choi et al. JASA (2010)

sail: Weak Heredity

Reparametrization

$$\boldsymbol{\tau}_j = \gamma_j(\beta_E \cdot \mathbf{1}_{m_j} + \boldsymbol{\theta}_j)$$

Model

$$Y = \beta_0 \cdot \mathbf{1} + \sum_{j=1}^p \boldsymbol{\Psi}_j \theta_j + \beta_E X_E + \sum_{j=1}^p \gamma_j (X_E \circ \boldsymbol{\Psi}_j) (\beta_E \cdot \mathbf{1}_{m_j} + \boldsymbol{\theta}_j) + \varepsilon$$

Objective Function

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}, \boldsymbol{\gamma}} \mathcal{L}(\boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

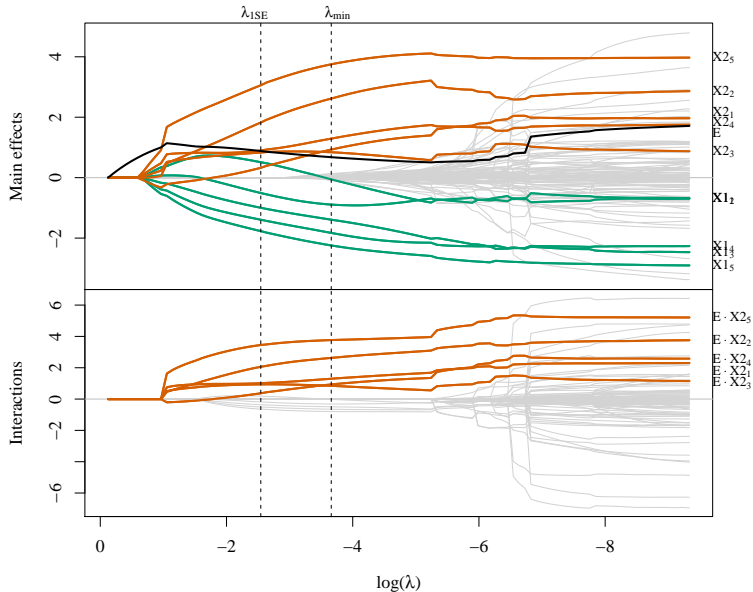
Toy example

- With a sample size of $n = 100$, we sample $p = 20$ covariates X_1, \dots, X_p independently from a $N(0, 1)$ distribution truncated to the interval $[0, 1]$.
- Data were generated from a model which follows the strong heredity principle, but where only one covariate, X_2 , is involved in an interaction with a binary exposure variable (E):

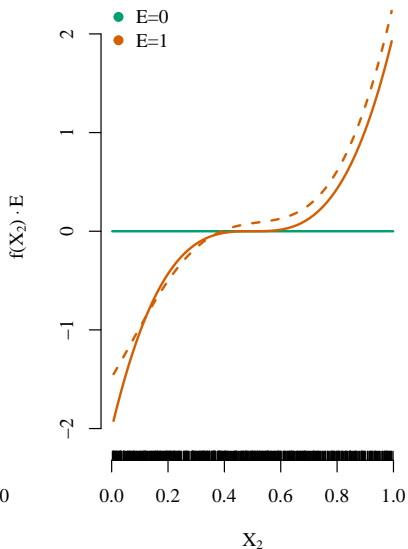
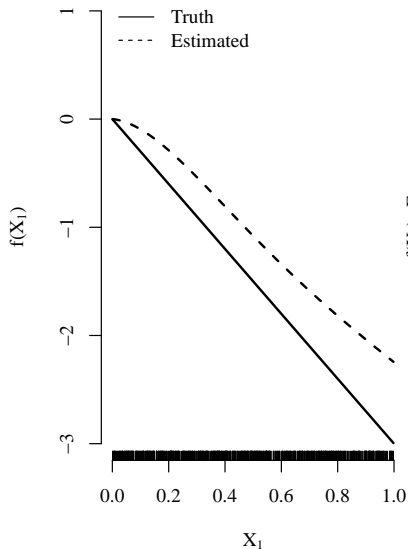
$$Y = f_1(X_1) + f_2(X_2) + 1.75E + 1.5E \cdot f_2(X_2) + \varepsilon.$$

- For illustration, function $f_1(\cdot)$ is assumed to be linear, whereas function $f_2(\cdot)$ is non-linear: $f_1(x) = -3x$, $f_2(x) = 2(2x - 1)^3$.

Toy example - Solution path



Toy example - Estimated effects



Block Relaxation (De Leeuw, 1994)

Algorithm 1: Block Relaxation Algorithm

Set the iteration counter $k \leftarrow 0$ and fix $\alpha \in (0, 1)$;

for each λ **do**

repeat

$$\gamma^{(k+1)} \leftarrow \operatorname{argmin}_{\gamma} Q_{\lambda} \left(\gamma, \beta_E^{(k)}, \boldsymbol{\theta}^{(k)} \right)$$

$$\boldsymbol{\theta}^{(k+1)} \leftarrow \operatorname{argmin}_{\boldsymbol{\theta}} Q_{\lambda} \left(\boldsymbol{\theta}, \beta_E^{(k)}, \gamma^{(k+1)} \right)$$

$$\beta_E^{(k+1)} \leftarrow \operatorname{argmin}_{\beta_E} Q_{\lambda} \left(\boldsymbol{\theta}^{(k+1)}, \beta_E, \gamma^{(k+1)} \right)$$

$$k \leftarrow k + 1$$

until convergence criterion is satisfied;

end

Implementation

Objective Function

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}, \boldsymbol{\gamma}} \mathcal{L}(Y; \boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

¹<https://cran.r-project.org/package=sail>

Implementation

Objective Function

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}, \gamma} \mathcal{L}(Y; \boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

Lasso problem

$$\operatorname{argmin}_{\gamma} \mathcal{L}(Y; \boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

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Group Lasso problem

$$\operatorname{argmin}_{\beta_E, \boldsymbol{\theta}} \mathcal{L}(Y; \boldsymbol{\Theta}) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

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Theorem 1

$$\hat{\Theta}_n = \operatorname{argmin}_{\beta_E, \theta, \gamma} \mathcal{L}(\Theta) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

$$\mathcal{A}_1 = \{j : \theta_j \neq 0, \beta_j \neq 0\}$$

$$\mathcal{A}_2 = \{k : \gamma_k \neq 0\}, \quad \mathcal{A} = \mathcal{A}_1 \cup \mathcal{A}_2$$

Under certain regularity conditions and the existence of a local minimizer $\hat{\Theta}_n$ that is \sqrt{n} -consistent

$$P\left(\hat{\Theta}_{\mathcal{A}^c} = 0\right) \rightarrow 1$$

Theorem 1

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Under certain regularity conditions and the existence of a local minimizer $\hat{\Theta}_n$ that is \sqrt{n} -consistent

$$P\left(\hat{\Theta}_{\mathcal{A}^c} = 0\right) \rightarrow 1$$

Theorem 1 shows that when the tuning parameters for the nonzero coefficients converge to 0 faster than $n^{-1/2}$ we can consistently remove the noise terms with probability tending to 1.

Asymptotic normality

Theorem 2

$$\hat{\Theta}_n = \operatorname{argmin}_{\beta_E, \theta, \gamma} \mathcal{L}(\Theta) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

Under certain regularity conditions, the component $\hat{\Theta}_{\mathcal{A}}$ of the local minimizer $\hat{\Theta}_n$ satisfies

$$\sqrt{n} \left(\hat{\Theta}_{\mathcal{A}} - \Theta_{\mathcal{A}} \right) \rightarrow_d \mathcal{N} \left(0, \mathbf{I}^{-1} \left(\Theta_{\mathcal{A}} \right) \right)$$

Theorem 2 shows that the `sail` estimates for nonzero coefficients in the true model have the same asymptotic distribution as they would have if the zero coefficients were known in advance.

Asymptotic normality

Theorem 2

$$\hat{\Theta}_n = \underset{\beta_E, \theta, \gamma}{\operatorname{argmin}} \quad \mathcal{L}(\Theta) + \lambda(1 - \alpha) \left(w_E |\beta_E| + \sum_{j=1}^p w_j \|\theta_j\|_2 \right) + \lambda\alpha \sum_{j=1}^p w_{jE} |\gamma_j|$$

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Theorem 2 shows that the `sail` estimates for nonzero coefficients in the true model have the same asymptotic distribution as they would have if the zero coefficients were known in advance.

Theorem 1 + 2 \rightarrow Oracle property (Fan and Li, 2001)

Simulation Scenarios

1. Truth obeys strong hierarchy (**right in our wheel house**):

$$Y = \sum_{j=1}^4 f_j(X_j) + \beta_E \cdot X_E + X_E \times (f_3(X_3) + f_4(X_4)) + \varepsilon$$

Simulation Scenarios

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2. Truth obeys weak hierarchy

Simulation Scenarios

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3. **Truth only has interactions**

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5. **Truth only has main effects**

Simulation Scenarios

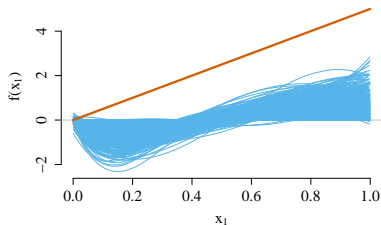
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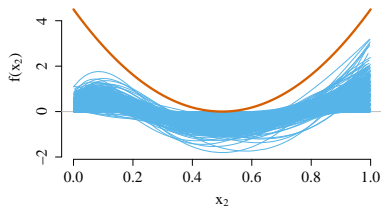
2. **Truth obeys weak hierarchy**
 3. **Truth only has interactions**
 4. **Truth is linear**
 5. **Truth only has main effects**
- $n_{train} = n_{tuning} = 200, n_{test} = 800, p = 1000, \beta_E = 1, SNR = 2$
 - $X_j \sim \text{truncnorm}(0, 1), j = 1, \dots, 1000, E \sim \text{truncnorm}(-1, 1)$
 - sail needs to estimate $1000 \times 5 \times 2 = 10\text{k}$ parameters

Scenario 1: Main Effects for 500 Simulations

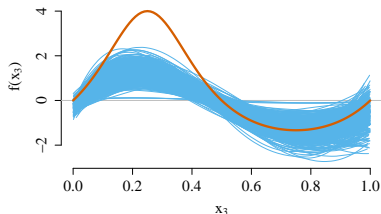
$$f(x_1) = 5x_1$$



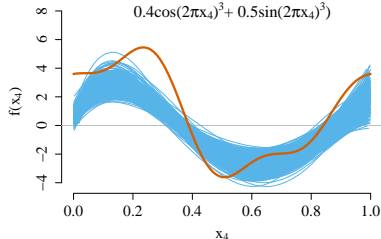
$$f(x_2) = 4.5(2x_2 - 1)^2$$



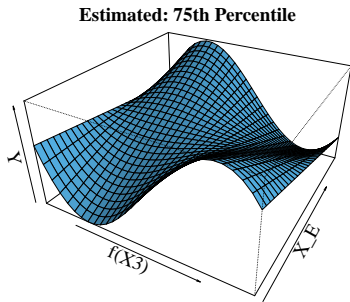
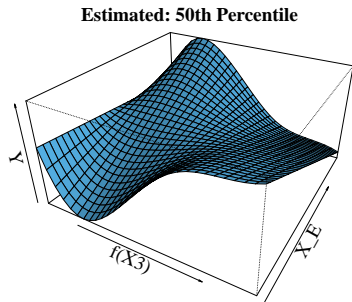
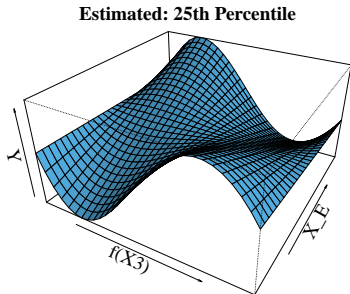
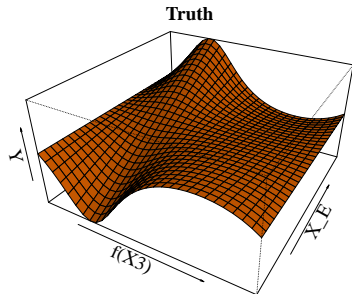
$$f(x_3) = \frac{4\sin(2\pi x_3)}{2 - \sin(2\pi x_3)}$$



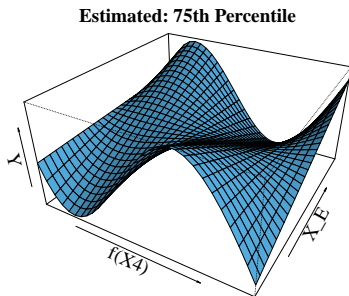
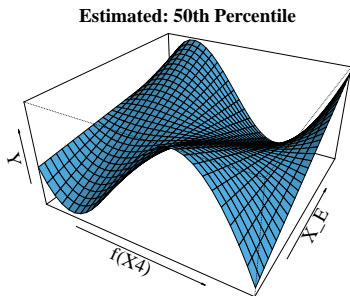
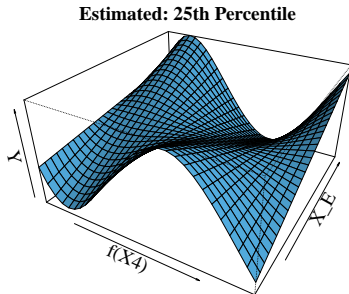
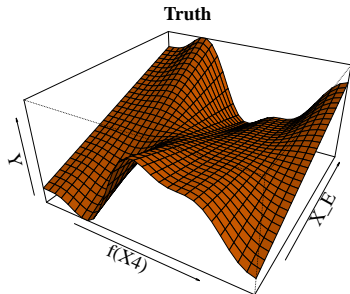
$$f(x_4) = 6(0.1\sin(2\pi x_4) + 0.2\cos(2\pi x_4) + 0.3\sin(2\pi x_4)^2 + 0.4\cos(2\pi x_4)^3 + 0.5\sin(2\pi x_4)^3)$$



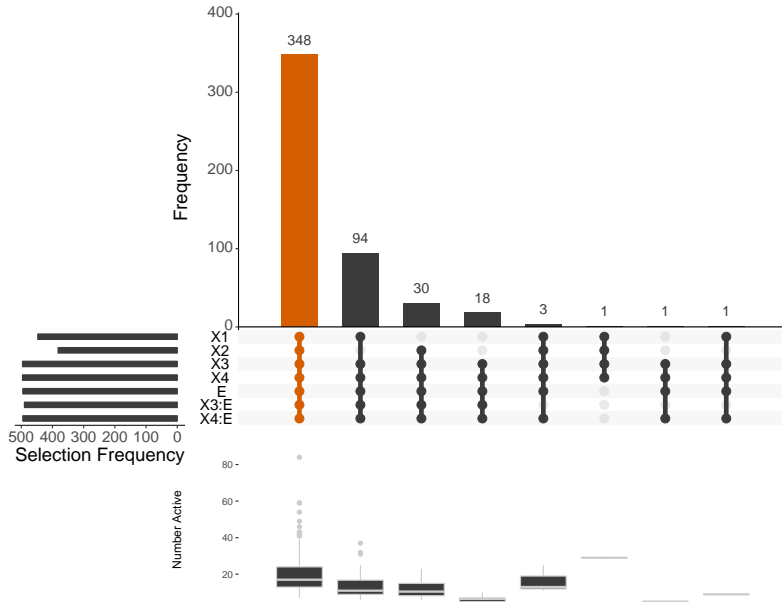
Scenario 1: Estimated Interaction Effects for $E \cdot f(X_3)$



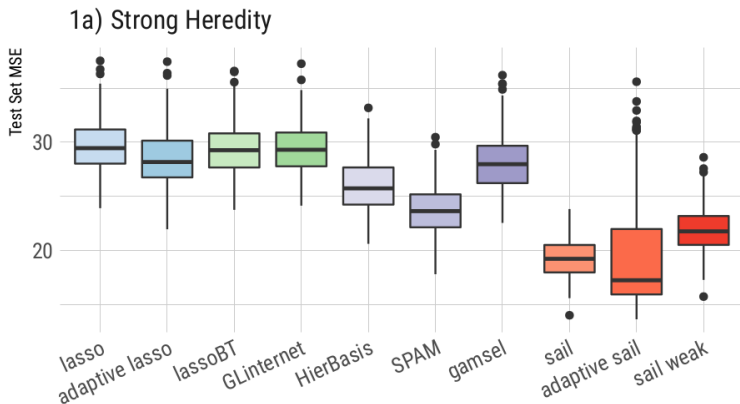
Scenario 1: Estimated Interaction Effects for $E \cdot f(X_4)$



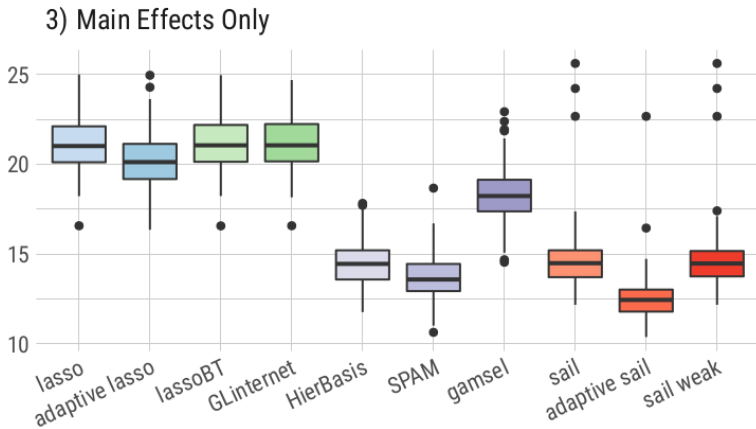
Right in Our Wheel House Simulation Results



Strong Heredity

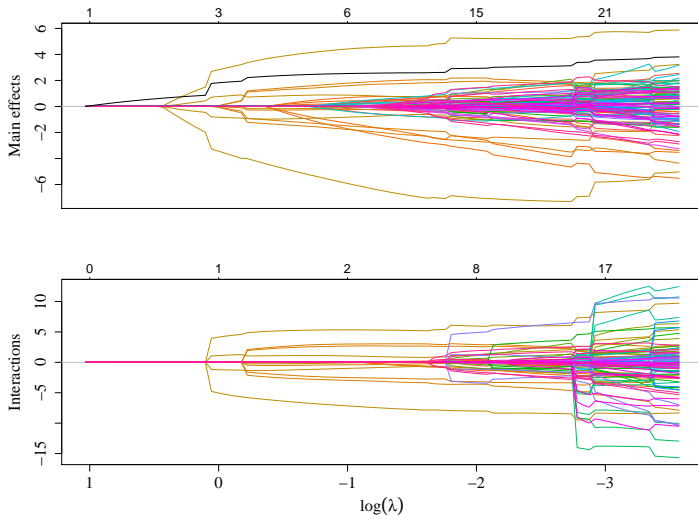


Main Effects Only



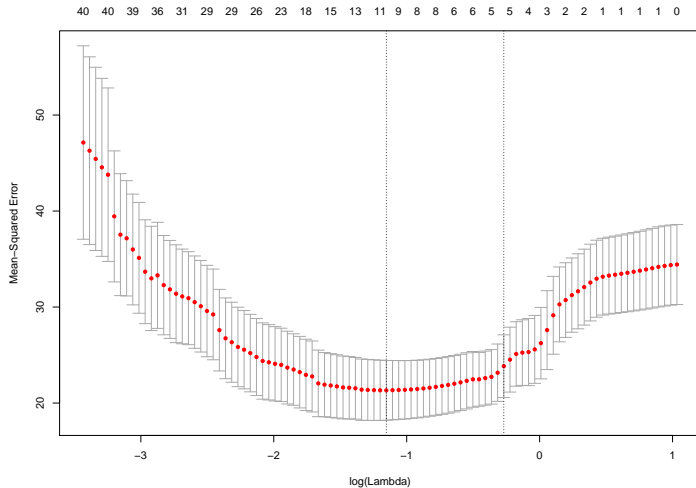
sail R package: Solution Path results

```
f.basis <- function(x) splines::bs(x, degree = 5)
fit <- sail(x, y, e, basis = f.basis)
plot(fit)
```



sail R package: Cross-validation results

```
sail::plot(cvfit)
```



Nurse Family Partnership Program

- Early intervention in young children has been shown to positively impact intellectual abilities.
- Genome-wide association studies (GWAS) suggest that 20% of the variance in educational attainment (years of education) may be accounted for by common genetic variation.
- An interesting query that arises is how the environment interacts with these genetics variants to predict measures of cognitive function.

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- In this context, individuals with a higher PRS have a propensity for higher educational attainment.
- The goal of this analysis was to determine if there was an interaction between genetic predisposition to educational attainment (X) and maternal participation in the NFP program (E) on child IQ at 4 years of age (Y).

Application of sail to NFP data

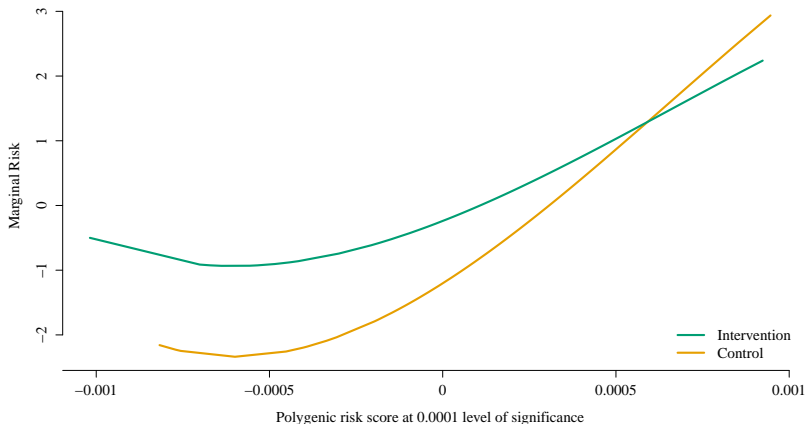


Fig.: The selected model, chosen via 10-fold cross-validation, contained three variables: the main effects for the intervention and the PRS for educational attainment using genetic variants significant at the 0.0001 level, as well as their interaction.

Strengths and Limitations

Strengths

- Non-linear environment interactions with strong heredity property in $p \gg N$
- `sail` allows for flexible modeling of input variables

Strengths and Limitations

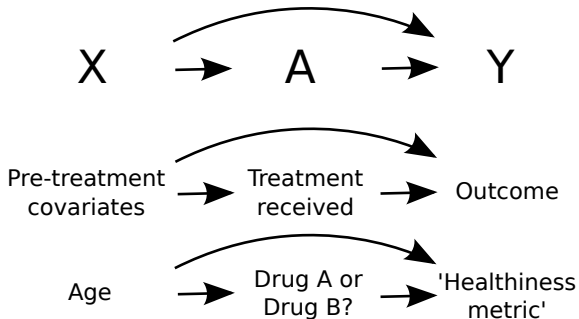
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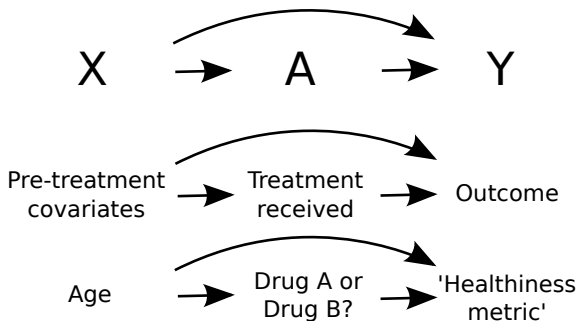
Limitations

- `sail` can currently only handle $E \cdot f(X)$ or $f(E) \cdot X$
- Does not allow for $f(X_1, E)$ or $f(X_1, X_2)$
- Memory footprint is an issue

Dynamic Treatment Regimes (DTRs)



Dynamic Treatment Regimes (DTRs)



$$\mathbb{E}[Y \mid \mathbf{X}, A; \psi, \beta] = \underbrace{\mathbf{X}\beta}_{\text{Impact of patient history in the absence of treatment}} + \underbrace{\psi_0 A + \psi A \mathbf{X}}_{\text{Impact of treatment on outcome}}$$

Extension of sail to DTRs



Cornell University

arXiv.org > stat > arXiv:2101.07359

Statistics > Methodology

[Submitted on 18 Jan 2021]

Variable Selection in Regression-based Estimation of Dynamic Treatment Regimes

Zeyu Bian, Erica EM Moodie, Susan M Shortreed, Sahir Bhatnagar

Dynamic treatment regimes (DTRs) consist of a sequence of decision rules, one per stage of intervention, that finds effective treatments for individual patients between treatment and a small number of covariates which are often chosen a priori. However, with increasingly large and complex data being collected, a driven approach of selecting these covariates might improve the estimated decision rules and simplify models to make them easier to interpret. We propose a method that has the strong heredity property, that is, an interaction term can be included in the model only if the corresponding main terms have also been selected. The newly proposed methods compare favorably with other variable selection approaches.

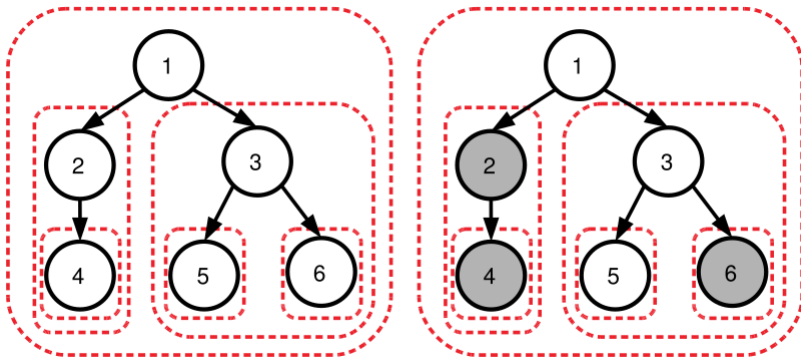
Subjects: **Methodology (stat.ME)**; Computation (stat.CO)

Cite as: [arXiv:2101.07359](https://arxiv.org/abs/2101.07359) [**stat.ME**]

(or [arXiv:2101.07359v1](https://arxiv.org/abs/2101.07359v1) [**stat.ME**] for this version)

¹*In revision at Biometrics.* <https://arxiv.org/abs/2101.07359>

Hierarchical Penalty Structure



¹Bach, Jenatton, Mairal and Obozinski (2011). Optimization with Sparsity-Inducing Penalties.

Bi-level selection

- Bi-level selection:

$$f(X_1) = \underbrace{\begin{bmatrix} X_{11} & \psi_{11}(X_{11}) & \psi_{12}(X_{12}) & \cdots & \psi_{11}(X_{15}) \\ & \vdots & \vdots & \cdots & \vdots \\ & \vdots & \vdots & \cdots & \vdots \\ X_{i1} & \psi_{11}(X_{i1}) & \psi_{12}(X_{i2}) & \cdots & \psi_{11}(X_{i5}) \\ & \vdots & \vdots & \cdots & \vdots \\ & \vdots & \vdots & \cdots & \vdots \\ X_{N1} & \psi_{11}(X_{N1}) & \psi_{12}(X_{N2}) & \cdots & \psi_{11}(X_{N5}) \end{bmatrix}}_{\Psi_1} \quad N \times 5 \quad \times \quad \underbrace{\begin{bmatrix} \beta_{\text{linear}} \\ \beta_{11} \\ \beta_{12} \\ \beta_{13} \\ \beta_{14} \\ \beta_{15} \end{bmatrix}}_{\theta_1} \quad 6 \times 1$$

Acknowledgements



Zeyu Bian, PhD (c)

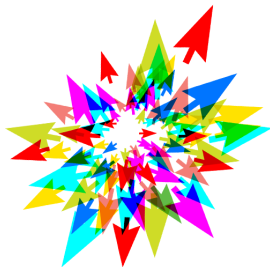


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- Celia Greenwood (Lady Davis Institute)
- Erica Moodie (McGill)
- Kieran O'Donnell (Yale)



compute | **calcul**
canada | canada



References

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2. **Bian Z**, Moodie EEM, Shortreed S, Bhatnagar SR (2021+). Variable Selection in Regression-based Estimation of Dynamic Treatment Regimes. <https://arxiv.org/abs/2101.07359>. *In revision at Biometrics*.
3. De Leeuw, J. (1994). Block-relaxation algorithms in statistics. In Information systems and data analysis (pp. 308-324). Springer Berlin Heidelberg.
4. Choi, N. H., Li, W., & Zhu, J. (2010). Variable selection with the strong heredity constraint and its oracle property. *Journal of the American Statistical Association*, 105(489), 354-364.
5. Chipman, H. (1996). Bayesian variable selection with related predictors. *Canadian Journal of Statistics*, 24(1), 17-36.

sahirbhatnagar.com

Session Info

```
R version 4.1.1 (2021-08-10)
Platform: x86_64-pc-linux-gnu (64-bit)
Running under: Pop!_OS 21.04
```

```
Matrix products: default
```

```
BLAS: /usr/lib/x86_64-linux-gnu/openblas-pthread/libblas.so.3
```

```
LAPACK: /usr/lib/x86_64-linux-gnu/openblas-pthread/libopenblas-p-r0.3.13.so
```

```
attached base packages:
```

```
[1] stats      graphics  grDevices  utils      datasets  methods    base
```

```
other attached packages:
```

```
[1] xtable_1.8-4      rpart.plot_3.1.0  rpart_4.1-15      data.table_1.14.2
[5] ISLR_1.2          ggplot2_3.3.5     knitr_1.36
```

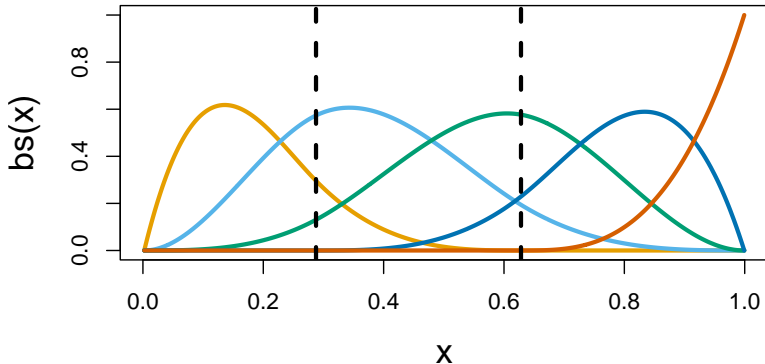
```
loaded via a namespace (and not attached):
```

```
[1] pillar_1.6.4      compiler_4.1.1    highr_0.9          tools_4.1.1
[5] digest_0.6.28     evaluate_0.14     lifecycle_1.0.1    tibble_3.1.5
[9] gtable_0.3.0      pkgconfig_2.0.3   rlang_0.4.12       DBI_1.1.1
[13] xfun_0.26         withr_2.4.2       dplyr_1.0.7        stringr_1.4.0
[17] generics_0.1.0    vctrs_0.3.8       grid_4.1.1         tidyselect_1.1.1
[21] glue_1.4.2        R6_2.5.1          fansi_0.5.0        pacman_0.5.1
[25] purrr_0.3.4       RSkittleBrewer_1.1 blob_1.2.1         magrittr_2.0.1
[29] scales_1.1.1      ellipsis_0.3.2    assertthat_0.2.1   colorspace_2.0-2
[33] utf8_1.2.2        stringi_1.7.5     munsell_0.5.0      crayon_1.4.1
```

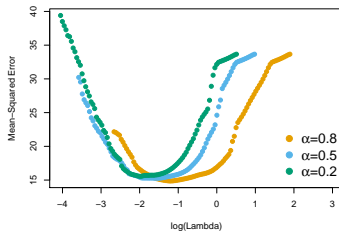
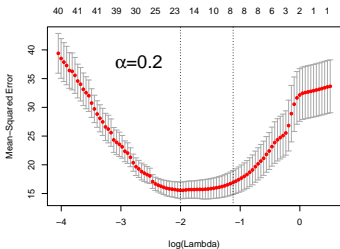
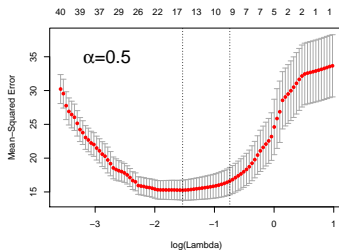
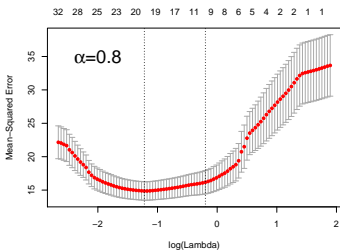

B-Spline Expansion

```
x <- truncnorm::rtruncnorm(1000, a = 0, b = 1)
B <- splines::bs(x, df = 5, degree=3, intercept = FALSE)
```

df=5, degree=3, inner.knots at c(33.33%, 66.66%) percentile



sail A Note on the Second Tuning Parameter results

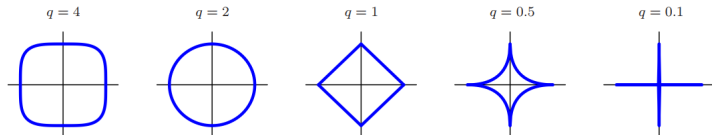


Why the L1 norm ?

- For a fixed real number $q \geq 0$ consider the criterion

$$\tilde{\beta} = \arg \min_{\beta} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^p |\beta_j|^q \right\}$$

- Why do we use the ℓ_1 norm? Why not use the $q = 2$ (Ridge) or any ℓ_q norm?



- $q = 1$ is the smallest value that yields a sparse solution **and** yields a **convex** problem \rightarrow scalable to high-dimensional data
- For $q < 1$ the constrained region is **nonconvex**

Linear Effects Simulation - Comparison

